## PRML chapter11

- 1. Markov Chain Monte Carlo
  - allows sampling from a large class of distributions, and which scales well with the dimensionality of the sample space
  - Metropolis algorithm
    - sample is accepted with probability

$$A\left(\mathbf{z}^{\star}, \mathbf{z}^{(\tau)}\right) = \min\left(1, \frac{\widetilde{p}\left(\mathbf{z}^{\star}\right)}{\widetilde{p}\left(\mathbf{z}^{(\tau)}\right)}\right)$$

- choose a random number u with uniform distribution over the unit interval (0,1)
- accepting the sample if  $A(\mathbf{z}^{\star}, \mathbf{z}^{(\tau)}) > u$
- a central goal in designing Markov chain Monte Carlo methods is to avoid random walk behaviour
- (a) Markov chains
  - Markov chain is the situation that we can approximate future from the given states
  - A first-order Markov chain is defined as

$$p\left(\mathbf{z}^{(m+1)}|\mathbf{z}^{(1)},\ldots,\mathbf{z}^{(m)}\right) = p\left(\mathbf{z}^{(m+1)}|\mathbf{z}^{(m)}\right)$$

• marginal probability is written as

$$p\left(\mathbf{z}^{(m+1)}\right) = \sum_{\mathbf{z}(m)} p\left(\mathbf{z}^{(m+1)} | \mathbf{z}^{(m)}\right) p\left(\mathbf{z}^{(m)}\right)$$

• A distribution is invariant when it satisfies

$$p^{\star}(\mathbf{z}) = \sum_{\mathbf{z}'} T\left(\mathbf{z}', \mathbf{z}\right) p^{\star}\left(\mathbf{z}'\right)$$

- (b) The Metropolis-Hastings algorithm
  - We first propose q(Z), then samples from it. We accept the sample when it sattisfies detailed balance
  - generalize metropolis algorithm by making

$$A_k\left(\mathbf{z}^{\star}, \mathbf{z}^{(\tau)}\right) = \min\left(1, \frac{\widetilde{p}\left(\mathbf{z}^{\star}\right) q_k\left(\mathbf{z}^{(\tau)} | \mathbf{z}^{\star}\right)}{\widetilde{p}\left(\mathbf{z}^{(\tau)}\right) q_k\left(\mathbf{z}^{\star} | \mathbf{z}^{(\tau)}\right)}\right)$$

• By using this we could proof the Metropolis algorithm samples from the required distribution

- 2. Gibbs Sampling
  - a special case of the Metropolis- Hastings algorithm.
  - First we initialize  $\{z_i : i = 1, \dots, M\}$
  - for each  $\tau$  : and for each j: sample

$$z_j^{(\tau+1)} \sim p\left(z_j | z_1^{(\tau+1)}, \dots, z_{j-1}^{(\tau+1)}, z_{j+1}^{(\tau)}, \dots, z_M^{(\tau)}\right)$$

- In order to gain the proper sampling, we need p(z) is invariant and Ergodicity
- we use "over-relaxation" to prevent it behaving like random walk.
- 3. Slice Sampling
  - Metropolis algorithm is sensitive to step size
  - The technique of slice sampling provides an adaptive step size that is automatically adjusted to match the characteristics of the distribution
- 4. The Hybrid Monte Carlo Algorithm
  - (a) Dynamical systems
    - $\bullet\,$  momentum variable

$$r_i = \frac{\mathrm{d}z_i}{\mathrm{d}\tau}$$

• under E(z) is potential energy

$$p(\mathbf{z}) = \frac{1}{Z_p} \exp(-E(\mathbf{z}))$$

• Physical energy is

$$K(\mathbf{r}) = \frac{1}{2} \|\mathbf{r}\|^2 = \frac{1}{2} \sum_{i} r_i^2$$

• The total energy of the system is

$$H(\mathbf{z}, \mathbf{r}) = E(\mathbf{z}) + K(\mathbf{r})$$

- During the evolution of this dynamical system, Hamilton function is constant
- Consider the joint distribution over phase space whose total energy is the Hamiltonian

$$p(\mathbf{z}, \mathbf{r}) = \frac{1}{Z_H} \exp(-H(\mathbf{z}, \mathbf{r}))$$

- leapfrog discretization
  - repeat following loop

$$\begin{aligned} \widehat{r}_i(\tau + \epsilon/2) &= \widehat{r}_i(\tau) - \frac{\epsilon}{2} \frac{\partial E}{\partial z_i}(\widehat{\mathbf{z}}(\tau)) \\ \widehat{z}_i(\tau + \epsilon) &= \widehat{z}_i(\tau) + \epsilon \widehat{r}_i(\tau + \epsilon/2) \\ \widehat{r}_i(\tau + \epsilon) &= \widehat{r}_i(\tau + \epsilon/2) - \frac{\epsilon}{2} \frac{\partial E}{\partial z_i}(\widehat{\mathbf{z}}(\tau + \epsilon)) \end{aligned}$$

- (b) Hybrid Monte Carlo
  - combine Hamilton dynamics and Metropolis algorithm
- (c) Estimating the Partition Function

- $\bullet$  in order to compare Bayes models, need to know the odds of  $Z_E$
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$$\begin{split} \frac{Z_E}{Z_G} &= \frac{\sum_{\mathbf{z}} \exp(-E(\mathbf{z}))}{\sum_{\mathbf{z}} \exp(-G(\mathbf{z}))} = \frac{\sum_{\mathbf{z}} \exp(-E(\mathbf{z}) + G(\mathbf{z})) \exp(-G(\mathbf{z}))}{\sum_{\mathbf{z}} \exp(-G(\mathbf{z}))} = E_{G(\mathbf{z})}[\exp(-E+G)] \\ &\simeq \frac{1}{L} \sum_{l} \exp\left(-E\left(\mathbf{z}^{(l)}\right) + G\left(\mathbf{z}^{(l)}\right)\right) \end{split}$$